

Régression en grande dimension et épistasie par blocs pour les études d'association

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Summary

GWAS and Block of linkage desequilibrium

- Genome Wide Association Studies
- Blocks of linkage desiquilibrium
- Hierachical Clustering with Adjacency Constraints
- How to improve?
- Some computation times
- 2 Epistasis
- 3 Method
 - The G-GEE modeling approach
 - Simulations
- Application
 - Ankylosing Spondylitis
 - First results

High-dimension in Genomics

'Omics'

• Genomics, Transcriptomics, Proteomics, Metabolomics, Epigenomics, Metagenomics

Large *p* small *n*

- n of the order of 10 to 10⁴ : Each statistical individual is a 'costly' experiment
- p of the order of 10 to 10^9 : Transcriptomics (10^4), Genomics (10^9) ...

Need for dimensionality reduction

- Selection (of variables)
- Projection in low dimensional subspace
- Clustering

High-dimension in Genomics

Technology evolution and Genomic

Genome

- Human Genome Project (1990-2003)
- 2002 launch of HapMap project (report in Nature 2005)
- 2008-2012 : the 1000 Genomes Project
- NGS (new generation sequencing)



Nature Biotechnology

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Single-Nucleotide Polymorphism Data

- 90 % of human genetic variation,
- In human genom, SNP with allelic frequency greater than 1 % are present every 300 base pairs (in average)
- 2 SNP among 3 substitute cytosine with thymine



Figure: SNP (wikipedia)

SNP Data



Genome-Wide Association Studies

GWAS characteristics :

• **Objective** : find associations between genetic markers $(SNP_{i,j} \in \{0, 1, 2\})$ and a phenotypic trait $(Y_i \in \{0, 1\})$ or $Y_i \in \mathbb{R}$



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• Generalized Linear Model

$$g(E[Y_i|x_i]) = \beta_0 + \sum_{j=1}^p \beta_j x_{ij} , i = 1, \dots, n$$

- *n* : number of individuals
- *p* : number of covariates
- Y_i : response for the individual i
- x_{.j} : observations for covariate j (coded in 0, 1 or 2)

SNP analysis

Differences between cases and controls at a specific SNP



Where is the missing heritability?

Missing Heritability

- gene-gene interaction,
- gene-environment interactions,
- rare variant effects (mutation present in less than 1% of the population),
- additivity of numerous common variants (not detected using univariate strategies)....

Data characteristics

- High dimension (p » n)
- Spatial structure
- Small effects

The LD measures

Linkage Desiquilibrium

- non-random association of alleles at two or more loci
- depends on the difference between observed allelic frequencies and those expected from a independent randomly distributed model.

Computation

Z_j the indicator of the presence of minor allele for SNP j.
Z_i ∼ B(p_i)

$$D(j,k) = p_{jk} - p_j p_k = E[Z_j Z_k] - p_j p_k = cov(Z_j, Z_k)$$
$$r^2(j,k) = corr(Z_j, Z_k)$$

ou

$$D'(j,k) = D(j,k)/Dmax(j,k)$$

How to estimate LD?

snp	vv	vV	VV	snn	V	V
uu	а	b	С	sip	V	
uU	d	е	f	<u> </u>	α	β
	~	h		U U	γ	δ
00	g	п	I			

Only the genotype data table is observed

- α , β , γ , δ are estimated
- a system of equations. e.g : $\alpha = 2a + b + d + pe$

with p the "probability" of the haplotype (uv, UV).

⇒ estimating p, then (α , β , γ , δ) and finally $D = p_{UV} - p_U p_V$.

The LD block structure

- the *r*² coefficients among the **50 first SNPs** of the Chromosome 22 (Dalmasso et al. 2008)
- LD structured in blocks



Hierachical Clustering with Adjacency Constraints



Block-Wise Approach using Linkage Disequilibrium (BALD)

- Hierarchical clustering of the SNPs with adjacency constraint and using the LD similarity.
- Estimation of the optimal number of groups using the Gap statistic (Tibshirani et. al., 2001).

- All coefficients outside the band "h" are null
- a $p \times h$ similarity matrix



 \Rightarrow a hierarchical clustering with adjacency constraint

A pseudocode

Data:
$$X \in \{0, 1, 2\}^{n \times p}$$
, Sim
 $C \leftarrow \{C_i = \{X_{.i}\}, i \in 1, ..., p\}$ /* clusters = singletons
*/;
 $D \leftarrow \{1 - Sim(X_{.i}, X_{.(i+1)}), i \in 1, ..., p - 1\}$;
for step = 1 to $p - 1$ do
 $i^* \leftarrow \operatorname{argmin}_{i \in \{1, ..., p - step\}} D(C_i, C_{i+1})$;
 $C \leftarrow C \setminus \{C_{i^*}, C_{i^*+1}\} \cup \{C_{i^*} \cup C_{i^*+1}\}$;
 $d_1 \leftarrow D(C_{i^*-1}, C_{i^*} \cup C_{i^*+1})$;
 $d_2 \leftarrow D(C_{i^*} \cup C_{i^*+1}, C_{i^*+2})$;
 $D \leftarrow D \setminus \{D(C_{i^*-1}, C_{i^*}), D(C_{i^*}, C_{i^*+1})\} \cup \{d_1, d_2\}$;
end

The Ward's distance

Ward Constrained Hierarchical Clustering

$$d(A,B) = rac{n_A n_B}{n_A + n_B} \left(rac{1}{n_A^2} S_{A,A} + rac{1}{n_B^2} S_{B,B} - rac{2}{n_A n_B} S_{A,B}
ight)$$



The pencils' trick : Calculating S_{AA} and S_{AB}





Assessing S_{AA} , S_{BB} and S_{AB} requires the calculation of sums of LD measures within *pencil-shaped areas* defined by :

- direction : right or left
- depth : hLoc
- end point : lim

 \Rightarrow Two arrays of sizes $p \times h$ for storing the pencils sums.

- All nodes are either less than or equal to each of its children.
- Uniquely represented by storing its level order traversal in an array. Given a position *i* :
 - $Parent(i) = \lfloor i/2 \rfloor$
 - Left(i) = 2i
 - Right(i) = 2i + 1



DeleteMin











Time complexity : O(log(p))

InsertHeap











Time complexity : O(log(p))

Data: An array A Result: A min-heap H for $i = \lfloor length(A)/2 \rfloor$ down to 1 do \mid PercolDown(A, i); end



Time complexity : $\mathcal{O}(plog(p))$

Time complexity of some operations

	findMin	insert	deleteMin
unordered array	$\mathcal{O}(p)$	$\mathcal{O}(1)$	$\mathcal{O}(p)$
binary heap	$\mathcal{O}(1)$	$\mathcal{O}(\log(p))$	$\mathcal{O}(log(p))$

cWard in seconds...



Figure: The mean computation time t versus the number of markers p for the cWard algorithm applied to randomly sampled SNP matrices. N = 100, h = 30 and t is averaged across 50 simulation runs.

Compared to a former implementation



Figure: The mean computation time t versus the number of markers p for the cWard algorithm and an implementation without heaps. t is averaged across 20 simulation runs.

Scalable Hierarchical Clustering with pencils and binary heap

To sum up :

- A $\mathcal{O}(p^2)$ algorithm does not scale for GWA studies.
- In the Ward distance written in a simple way.
- Space complexity of $\mathcal{O}(ph)$ by using the pencils' trick.
- Time complexity of :



 $\mathcal{O}(plog(p))$

building the heap and insert/delete heaps' operations within the loop

Ongoing work :

Currently implemented with a genotype matrix as input.
 ⇒ can be generalized to any band similarity matrix.

What is the appropriate scale for association analysis?



AUC = f(Nb cluster) on real dataset (300 cut level)

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Definition

Interaction of alleles effects from different markers

Existing methods

- mainly SNP x SNP
- some at the block (gene) scale

Advantages of gene (or block) scale approaches

- results biologically interpretable
- genetic effects may be easier to detect
- reduce the number of variables

Epistasis - Gene scale methods

Existing gene scale methods :

Two or few genes

- PCA + logistic regression (He et al. 2011, Li et al. 2009, Zhang et al. 2008)
- PLS + logistic regression (Wang T et al. 2009)

For a larger number of genes

- PCA + LASSO (D'Angelo et al. 2009)
- PCA + pathway-guided penalized regression (Wang X et al. 2014)

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Objectives : To develop a new gene scale method

- considers interaction variables,
- takes into account the group structure,
- is applicable with many groups (genes)

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Group modeling approach



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model :

$$g(E[\boldsymbol{y}|\boldsymbol{X}]) = \underbrace{\sum_{g} \sum_{p_g} \beta_{g,p_g} \boldsymbol{X}_{g,p_g}}_{\text{Main effects}}$$

$$\boldsymbol{\beta} = \left(\underbrace{\beta_{1,1}, \beta_{1,2}, \cdots, \beta_{1,p_1}}, \cdots, \underbrace{\beta_{G,1}, \cdots, \beta_{G,p_G}}, \underbrace{\beta_{G,2}, \cdots, \beta_{G,p_$$

Group modeling approach



model :

$$g(E[y|X]) = \underbrace{\sum_{g} \sum_{p_g} \beta_{g,p_g} X_{g,p_g}}_{\text{Main effects}} + \underbrace{\sum_{r,s} \gamma_{r,s} Z_{r,s}}_{\text{Interaction effects}}$$

$$\boldsymbol{\beta} = \left(\underbrace{\beta_{1,1}, \beta_{1,2}, \cdots, \beta_{1,p_1}}_{35/48}, \cdots, \underbrace{\beta_{G,1}, \cdots, \beta_{G,p_G}}_{35/48}\right)^T \qquad \boldsymbol{\gamma} = \left(\begin{array}{ccc} \boldsymbol{\gamma}_{12}, \cdots, \begin{array}{c} \boldsymbol{\gamma}_{1G} & \cdots, \boldsymbol{\gamma}_{(G-1)G} \\ \end{array}\right)^{35/48}$$

Interaction variables construction : Gene-Gene Eigen Epistasis (G-GEE)

$$Z^{rs} = f_{u}(X^{r}, X^{s}) \text{ the interaction between genes } r, s.$$
$$\hat{u} = \arg \max_{u, ||u||=1} cov^{2}(y, f_{u}(X^{r}, X^{s}))$$

We set :
$$f_{\boldsymbol{u}}(\boldsymbol{X}^r, \boldsymbol{X}^s) = \boldsymbol{W}^{rs} \boldsymbol{u}$$
 with
 $\boldsymbol{W}^{rs} = \{X_{ij}^r X_{ik}^s\}_{i=1\cdots n}^{j=1\cdots, p_r; k=1,\cdots, p_s}$

$$\max_{\boldsymbol{u},\|\boldsymbol{u}\|=1} ||c\hat{o}v[\boldsymbol{W}^{rs}\boldsymbol{u},\boldsymbol{y}]||^2 = \max_{\boldsymbol{u},\|\boldsymbol{u}\|=1} \boldsymbol{u}^T \boldsymbol{W}^{rsT} \boldsymbol{y} \boldsymbol{y}^T \boldsymbol{W}^{rs} \boldsymbol{u}$$

u : eigen vector associated to the largest eigenvalue of $W^{rs \, T} y y^{\, T} W^{rs}$

$$u = W^{rsT}y$$

Coefficients estimation (linear model)

Group LASSO regression

$$\begin{pmatrix} \hat{\beta}, \hat{\gamma} \end{pmatrix} = \underset{\beta, \gamma}{\operatorname{argmin}} \sum_{i} (y_{i} - \boldsymbol{X}_{i}\beta - \boldsymbol{Z}_{i}\gamma)^{2} + \\ \lambda \left(\sum_{g} \sqrt{p_{g}} ||\beta^{g}||_{2} + \sum_{rs} \sqrt{p_{r}p_{s}} ||\gamma^{rs}||_{2} \right)$$

Limits of the groupLASSO regression :

•
$$P(S^* \subset \hat{S}) \xrightarrow[n \to +\infty]{} 1 \text{ but } |\hat{S}| >> |S^*|$$

• Difficult to compute p-value or confidence interval

Coefficients estimation (linear model)

Group LASSO regression

$$(\hat{\boldsymbol{\beta}}, \hat{\boldsymbol{\gamma}}) = \underset{\boldsymbol{\beta}, \boldsymbol{\gamma}}{\operatorname{argmin}} \sum_{i} (y_{i} - \boldsymbol{X}_{i}\boldsymbol{\beta} - \boldsymbol{Z}_{i}\boldsymbol{\gamma})^{2} + \\ \lambda \left(\sum_{g} \sqrt{p_{g}} ||\boldsymbol{\beta}^{g}||_{2} + \sum_{rs} \sqrt{p_{r}p_{s}} ||\boldsymbol{\gamma}^{rs}||_{2} \right)$$

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Adaptive-Ridge Cleaning Becu JM, 2015

- Use of a specific penalty for group LASSO
- Permutation test based on Fisher test approach for each group

Simulations design

Genotype :

From a real data set composed of 763 individuals and 63340 SNPs structured in 7216 genes.

Continuous phenotype simulated under two different schemes :

→ from Wang X et al., 2014 :

$$Y_{i} = \beta_{0} + \sum_{g} \beta_{g} \left(\sum_{k \in \mathcal{C}} X_{ik}^{g} \right) + \sum_{rs} \gamma_{rs} \left(\sum_{(j,k) \in \mathcal{C}^{2}} X_{ik}^{r} X_{ik}^{s} \right) + \epsilon_{i} \quad (1)$$

→ PCA model :

$$Y_{i} = \beta_{0} + \sum_{g} \beta_{g} \left(\sum_{k \in \mathcal{C}} X_{ik}^{g} \right) + \sum_{rs} \gamma_{rs} C_{i1}^{r} C_{i1}^{s} + \epsilon_{i}.$$
(2)

Simulations design

Scenarios :

For each setting we consider 6 genes.

→Five settings :

- same genes for main and interaction effects,
- different genes for main and interaction effects,
- only one interaction effect,
- only two main effects,
- no effects.

Simulations results - Interactions power



40 / 48

Simulations results - Discoveries matrix



Simulations results - $R^2 = 0.2$



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Ankylosing Spondylitis

Chronic inflammatory disease of the axial skeleton

Epidemiology :

- Age at first symptoms : 20 30 years
- Sexe : predominance for men (sex ratio 2M :1W)
- Prevalence : depend of populations (0.1% 1.4%)

Right etiology unknown :

- Environmental factors?
- Genetic factors?
 → Importance of HLA complex

HLA complex :

- Localized on chromosome 6
- Regroup about 200 genes
- Coding the immunity system
- Antigen HLA-B27 : associated to SPA

→ Effect from other gene in HLA group?

Known genes

Table I Summary of ankylosing spondylitis-susceptibility genes

identified by genome-wide association studies		IL6R	Interleukin 6 receptor		
RUNX3 IL23R	Runt-related transcription factor 3	FCGR2A	Fc fragment of immunoglobulin G, low-affinity IIa, receptor (CD32)		
ILI 2RB2	Interleukin 12 receptor, B2	UBE2E3	Ubiquitin-conjugating enzyme E2E 3		
GRP25	G-protein-coupled receptor 25	GPR35	G-protein-coupled receptor 35		
KIF21B	Kinesin family member 21B	NKX2-3	NK2 homeobox 3		
PTGER4	Prostaglandin E receptor 4 (subtype EP.)	ZMIZI	Zinc finger, MIZ type-containing I		
ERAPI	Endoplasmic reticulum aminopeptidase I	SH2B3	Src homology 2B adaptor protein 3		
ERAP2	Endoplasmic reticulum aminopeptidase 2	GPR65	G-protein-coupled receptor 65		
LNPEP	Leucyl/cystinyl aminopeptidase	IL27	Interleukin 27		
ILI 2B	Interleukin I2B	SULTIAI	Sulfotransferase family cytosolic 1A		
CARD9	Caspase recruitment-domain family member 9	TYK2	Tyrosine kinase 2		
LTβR	Lymphotoxin β-receptor (TNFR superfamily, member 3)	ICOSLG	Inducible T-cell costimulator ligand		
TNFRSFIA	Tumor-necrosis factor-receptor superfamily member IA	EOMES	Eomesodermin		
NPEPPS	Aminopeptidase puromycin-sensitive	IL7R	Interleukin 7 receptor		
TBKBPI	TNFR-associated factor family member-associated	BACH2	BTB and CNC homology 1, basic leucine-zipper		
	nuclear factor-xB-binding kinase 1-binding protein		transcription-factor 2		
TBX21	TBX2/ T-box 21		Abbreviation: CD, classification determinant.		

Tsui et al., 2014 : The genetic basis of ankylosing spondylitis : new insights into disease pathogenesis, The Application of Clinical Genetics :7 105-115

\rightarrow 29 susceptibility genes identified by GWAS

	Significant results
G-GEE	HLA-B × SULT1A1
	IL23R x ERAP2
PLS	HLA-B
	EOMES × BACH2
PCA	HLA-B

Conclusions and perspectives

The G-GEE method

- Takes into account the gene structure of data
- Can be applied on a large number of genes
- Uses a specific interaction modeling approach

Ankylosing Spondylitis

- Identification of potential interactions to discuss with medical doctors
- HLA-B effect

Perspectives

- Explore new f_u(X^r_i, X^s_i)
- Definition Additional simulations on larger data set
- New applications on other data sets

Thank you for your attention !



Simulations results - Main effect power



Group LASSO regression

$$\hat{\boldsymbol{\theta}} = (\hat{\boldsymbol{\beta}}, \hat{\boldsymbol{\gamma}}) = \operatorname*{argmin}_{\boldsymbol{\beta}, \boldsymbol{\gamma}} \left(\sum_{i} - L(y_i; \boldsymbol{X}_i \boldsymbol{\beta} + \boldsymbol{Z}_i \boldsymbol{\gamma}) + \lambda \left[\sum_{g} \sqrt{p_g} ||\boldsymbol{\beta}^g||_2 + \sum_{rs} \sqrt{p_r p_s} \right] \right)$$

Execution median time (seconds)



Methods - GGEE - PCA - PLS

Simulations design

Genotype : $\mathbf{X}_i \sim \mathcal{N}_{\rho}(\mathbf{0}, \mathbf{\Sigma})$ with $\mathbf{\Sigma}$ a block diagonal correlation matrix $(\rho = 0.8 \text{ for two SNPs in the same gene})$

 $\textit{MAF}_{j} \sim \mathcal{U}[0.05, 0.5]$ with $\textit{MAF}_{j} = 0.2$ if j causal SNP

Scenarios :

We consider 600 subjects and 6 SNPs by gene

→First scenario on 6 genes, two settings :

- same genes for main and interaction effects,
- different genes for main and interaction effects.

 \rightarrow Second scenario on 25 genes, one setting :

• different genes for main and interaction effects.

Simulations results - Interactions power; First scenario on 6 genes

Wang model

PCA model



Simulations results - $R^2 = 0.2$; First scenario on 6 genes



Wang model

PCA model



- → Main effects :
 - gene 1
 - gene 2
- → Interaction

- → Main effects : gene 1 gene 2
 - . . .

Simulations results - Second scenario on 25 genes



→ Main effects :

gene 1 gene 2

- → Interaction effects : gene 3 x gene 4
 - gene 5 x gene 6 gene 7 x gene 8
 - gene 9 x gene 10

Figure: Wang X et al. model, $R^2 = 0.7$

specific penalty for group LASSO : $\frac{\lambda}{\sqrt{|k(j)|\sum_{m \in k(j)} \hat{\theta}_m^2}}$