



Geostatistics for point processes

Predicting the intensity of ecological point process data

Edith Gabriel^{1,2} & Joël Chadœuf²

¹Laboratoire de Mathématiques d'Avignon ; ²Unité BioSP, INRA

JOURNÉES MAS DE LA SMAI
Grenoble, Août 2016

Spatial statistics

Study of stochastic phenomena $Y = \{Y(x), x \in D \subset \mathbb{R}^d\}$,
where Y is a random variable measured at spatial locations x_1, \dots, x_n .

Three main components:

- ★ **Geostatistics:**
spatial data indexed over a continuous space
- ★ Lattice data: spatial data indexed over a lattice of points
- ★ **Spatial point patterns:**
pertaining to the location of *events* of interest



Motivations

Predicting the local intensity

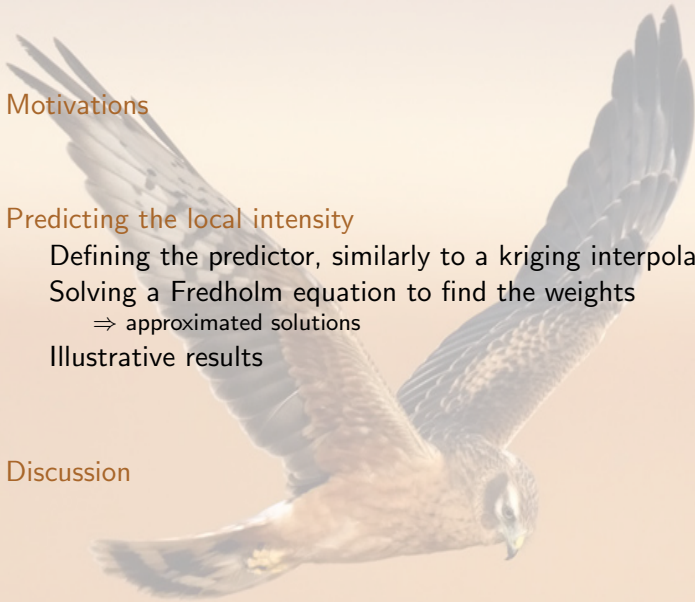
Defining the predictor, similarly to a kriging interpolator

Solving a Fredholm equation to find the weights

⇒ approximated solutions

Illustrative results

Discussion



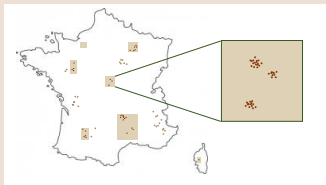
Motivating example



How to know the spatial distribution of a bird species at a national scale from observations made in a limited number of windows of few hectares?

The issue

How to extensively map the intensity of a point process in a large window when observation methods are available at a much smaller scale only?



⇒ The intensity of the process must be predicted from data issued out of samples spread over the window of interest.

Context

Let Φ a point process assumed to be

- stationary and isotropic,

$$\lambda = \frac{\mathbb{E}[\Phi(S_{obs})]}{\nu(S_{obs})} ; g(r) = \frac{1}{2\pi r} \frac{\partial K^*(r)}{\partial r}$$

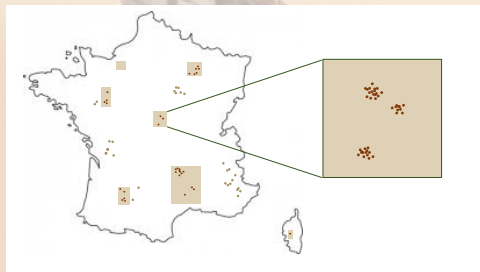
with $K^*(r) = \frac{1}{\lambda} \mathbb{E}[\Phi(b(0, r)) - 1 | 0 \in \Phi]$.

- observed in S_{obs} ,
- driven by a stationary random field, Z .

Our aim

Local intensity

We call *local* intensity of the point process Φ , its intensity given the random field, Z : $\lambda(x|Z)$.



Window of interest:

$$\begin{aligned} S &= S_{obs} \cup S_{unobs} \\ &= (\cup \square) \cup (\cup \square) \end{aligned}$$

$$\Phi = \{\circ, \bullet\}; \Phi_{S_{obs}} = \{\bullet\}$$

Our aim

To predict the local intensity in an unobserved window S_{unobs} .

Example

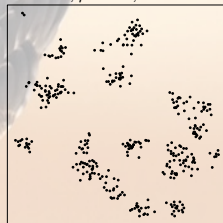
Thomas process:

- κ : intensity of the Poisson process parents, Z ,
- μ : mean number of offsprings per parent,
- σ : standard deviation of Gaussian displacement.

This process is stationary with intensity $\lambda = \kappa\mu$.

The local intensity corresponds to the intensity of the inhomogeneous Poisson process of offsprings, i.e. the intensity conditional to the parent process Z .

$$\kappa = 15, \mu = 15, \sigma = 0.025$$



★ More generally, we consider any process driven by a stationary random field ★

Existing solutions

- From the reconstruction of the process
 - Reconstruction method based on the 1st and 2^d-order characteristics of Φ (see e.g. [Tscheschel & Stoyan, 2006](#)).
 - Get the intensity by kernel smoothing.

A simulation-based method \Rightarrow long computation times.

- Intensity driven by a stationary random field
 - [Diggle *et al.* \(2007, 2013\)](#): Bayesian framework
 - [Monestiez *et al.* \(2006, 2013\)](#): Close to classical geostatistics.

Models constrained within the class of Cox processes.

- [van Lieshout and Baddeley \(2001\)](#).

Our alternative approach

We want to predict the local intensity $\lambda(x|Z)$

- outside the observation window,
- without precisely knowing the underlying point process
⇒ we only consider the 1st and 2^d-order characteristics,
- in a reasonable time.

We define an unbiased linear predictor

- which minimizes the error prediction variance (as in the geostatistical concept).
- whose weights depend on the structure of the point process.

Our predictor

Proposition

The predictor $\hat{\lambda}(x_o|Z) = \sum_{x \in \Phi \cap S_{obs}} w(x)$ is the BLUP of $\lambda(x_o|Z)$.

The weights, $w(x)$, are solution of the Fredholm equation of the 2^d kind:

$$\begin{aligned} w(x) + \lambda \int_{S_{obs}} w(y) (g(x-y) - 1) dy - \frac{1}{\nu(S_{obs})} \left[1 + \lambda \int_{S_{obs}^2} w(y) (g(x-y) - 1) dx dy \right] \\ = \lambda (g(x_o - x) - 1) - \frac{\lambda}{\nu(S_{obs})} \int_{S_{obs}} (g(x_o - x) - 1) dx \end{aligned}$$

and satisfy $\int_{S_{obs}} w(x) dx = 1$.

Solving the Fredholm equation

Any existing solution already considered in the literature can be used!

Our aim is to map the local intensity in a given window
 \Rightarrow access to fast solutions.

Several approximations can be used to solve the Fredholm equation.

The weights $w(x)$ can be defined as

- step functions \rightsquigarrow direct solution,
- linear combination of known basis functions, e.g. splines
 \rightsquigarrow continuous approximation.
- ...

Here, we illustrate the ones with the less heavy calculations and implementation.

Step functions

Let consider the following partition of S_{obs} : $S_{obs} = \cup_{j=1}^n B_j$, with

B : elementary square centered at 0,

$B_j = B \oplus c_j$: elementary square centered at c_j ,

$B_k \cap B_j = \emptyset$,

n : number of grid cell centers lying in S_{obs} .



For $w(x) = \sum_{j=1}^n w_j \frac{\mathbf{1}_{\{x \in B_j\}}}{\nu(B)}$, we get $\hat{\lambda}(x_o | Z) = \sum_{j=1}^n w_j \frac{\Phi(B_j)}{\nu(B)}$,

with $w = (w_1, \dots, w_n) = C^{-1}C_o + \frac{1 - \mathbf{1}^T C^{-1} C_o}{\mathbf{1}^T C^{-1} \mathbf{1}} C^{-1} \mathbf{1}$, where

- $C = \lambda \nu(B) \mathbb{I} + \lambda^2 \nu^2(B)(G - 1)$: covariance matrix
with $G = \{g_{ij}\}_{i,j=1,\dots,n}$, $g_{ij} = \frac{1}{\nu^2(B)} \int_{B \times B} g(c_i - c_j + u - v) du dv$,
and \mathbb{I} the $n \times n$ -identity matrix.
- $C_o = \lambda \nu(B) \mathbb{I}_{x_o} + \lambda^2 \nu^2(B)(G_o - 1)$: covariance vector
with \mathbb{I}_{x_o} the n -vector with zero values and one term equals to one where $x_o = c_i$,
and $G_o = \{g_{io}\}_{i=1,\dots,n}$.

Step functions: variance of the predictor

We consider the Neuman series to invert the covariance matrix, $C = \lambda\nu(B)\mathbf{1} + \lambda^2\nu^2(B)(G - 1)$, when $\lambda\nu(B) \rightarrow 0$:

$$C^{-1} = \frac{1}{\lambda\nu(B)} [\mathbf{1} + \lambda\nu(B)J_\lambda],$$

where a generic element of the matrix J_λ is given by

$$J_\lambda[i, j] = \sum_{k=1}^{\infty} (-1)^k \lambda^{k-1} (g(x_i, x_{l_1}) - 1) (g(x_{l_{k-1}}, x_j) - 1) \\ \times \int_{S_{obs}^{k-1}} \prod_{m=1}^{k-2} (g(x_{l_m}, x_{l_{m+1}}) - 1) dx_{l_1} \dots dx_{l_{k-1}}.$$

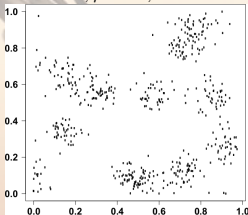
This leads to

$$\text{Var}(\widehat{\lambda}(x_o|Z)) = \lambda^3\nu^2(B)(G_o - 1)^T(G_o - 1) + \lambda^4\nu^3(B)(G_o - 1)^T J_\lambda(G_o - 1) \\ + \frac{1 - [\lambda\nu(B)\mathbf{1}^T(G_o - 1) + \lambda^2\nu^2(B)\mathbf{1}^T J_\lambda(G_o - 1)]^2}{\frac{\nu(S_{obs})}{\lambda} + \nu^2(B)\mathbf{1}^T J_\lambda \mathbf{1}}.$$

Step functions: illustrative results about prediction

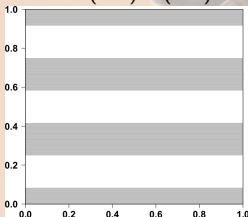
Simulated Thomas process

$$\kappa = 10, \mu = 50, \sigma = 0.05$$

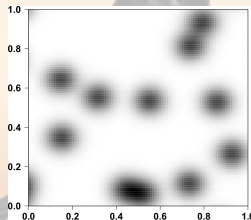


$$S = S_{obs} \cup S_{unobs}$$

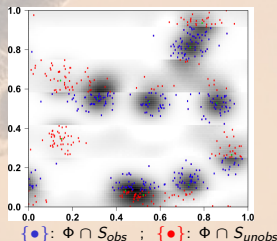
$$= (\cup \square) \cup (\cup \square)$$



Theoretical local intensity



Prediction within S_{unobs}



Spline basis

Let consider that the weights of $\hat{\lambda}(x_o|Z) = \sum_{x \in \Phi \cap S_{obs}} w(x)$ are defined as a degree d spline curve:

$$w(x) = \sum_{i=1}^k h_{i,d}(x),$$

where $h_{i,d}$ denotes the i th B -spline of order d .

A simplistic toy example in \mathbb{R} :

- $S_{obs} = [0, L] \subset [0, L'] = S$
- Linear spline defined from equally-spaced knots x_i :

$$w(x) = \begin{cases} a_0 + b_0x, & x \in \Delta_0 = [x_0, x_1) = [0, \frac{L}{k}), \\ a_1 + b_1x, & x \in \Delta_1 = [x_1, x_2) = [\frac{L}{k}, \frac{2L}{k}), \\ \vdots & \\ a_{k-1} + b_{k-1}x, & x \in \Delta_{k-1} = [x_{k-1}, x_k) = [\frac{(k-1)L}{k}, L), \end{cases}$$

$$= (a_i + b_i(x - x_i)) \mathbf{1}_{\{x \in \Delta_i\}}$$

Spline basis

From the continuity property and the constraint $\int_{S_{obs}} w(x) dx = 1$:

$$w(x) = \frac{1}{L} - \sum_{j=0}^{k-1} b_j P_j(x),$$

with $P_j(x) = \sum_{i=0}^{k-1} \left(\frac{1/2-k+j}{k^2} - \mathbf{1}_{\{j < i\}} - (x - \frac{il}{k}) \mathbf{1}_{\{i = j\}} \right) \mathbf{1}_{\{x \in \Delta_i\}}$

The Fredholm equation becomes

$$\begin{aligned} \sum_{j=0}^{k-1} b_j \left[P_j(x) + \lambda \int_L P_j(y) (g(x-y) - 1) dy - \frac{1}{L} \int_{L^2} P_j(y) (g(x-y) - 1) dx dy \right] \\ = \frac{\lambda}{L} \int_L (g(x-y) - 1) dy - \frac{1}{L^2} \int_{L^2} (g(x-y) - 1) dx dy - \lambda (g(x_0 - x) - 1) \\ + \frac{1}{L} \int_L (g(x_0 - x) - 1) dx \end{aligned}$$

i.e. of the form $\sum_{j=0}^{k-1} b_j A_j(x) = Q(x)$,

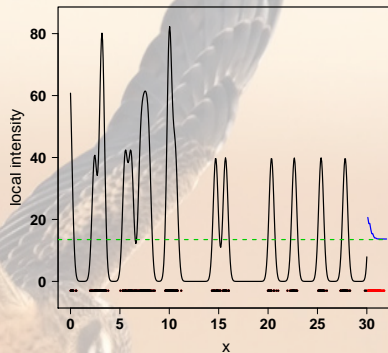
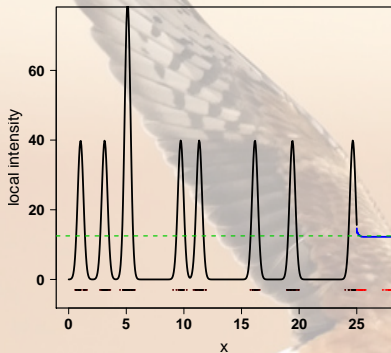
Then, $(b_0, \dots, b_{k-1}) = b$ is obtained from m control points and satisfy

$$b = (X^T X)^{-1} X^T Y,$$

with $X = (A_j(x_l))_{l=1, \dots, m}$ and $Y = (Q(x_l))_{l=1, \dots, m}$.

Spline basis: illustrative results

Thomas process in 1D ($\kappa = 0.5, \mu = 25, \sigma = 0.25$)



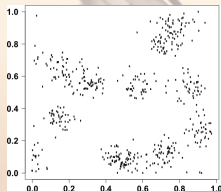
— Theoretical local intensity on S_{Obs} ; — Predicted values ; — Intensity of Φ

$$\{\bullet\} = \Phi_{S_{Obs}} ; \{\bullet\} = \Phi_{S_{Unobs}}$$

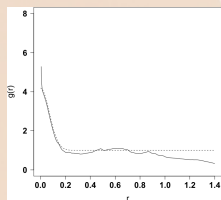
In practice: g must be estimated

Simulated Thomas process

$$\kappa = 10, \mu = 50, \sigma = 0.05$$

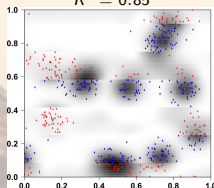


Estimated pcf



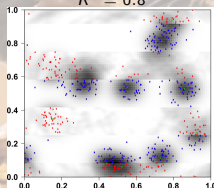
Prediction within S_{unobs}
(with the theoretical pcf)

$$R^2 = 0.85$$



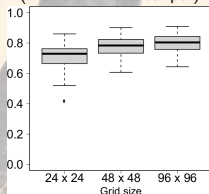
(with the estimated pcf)

$$R^2 = 0.8$$

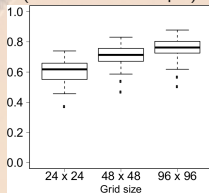


$\{\bullet\}$: $\Phi \cap S_{obs}$; $\{\bullet\}$: $\Phi \cap S_{unobs}$

R^2 in linear regression
of predicted and theoretical values
(with the theoretical pcf)



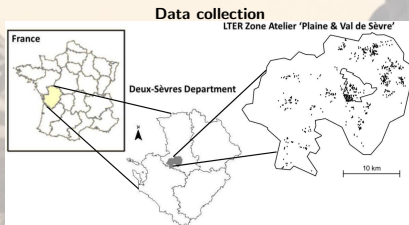
(with the estimated pcf)



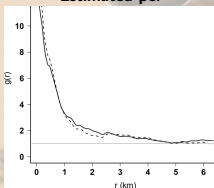
Application



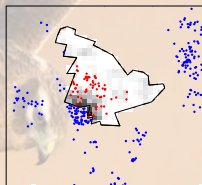
Montagu's Harriers' nest locations



Estimated pcf

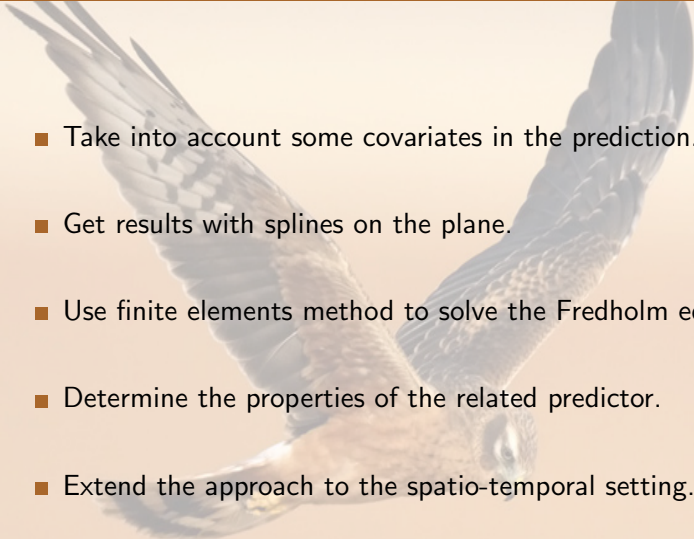


Prediction



$$\{\bullet\} = \Phi_{S_{obs}} ; \{\bullet\} = \Phi_{S_{unobs}}$$

Work in progress

- 
- Take into account some covariates in the prediction.
 - Get results with splines on the plane.
 - Use finite elements method to solve the Fredholm equation.
 - Determine the properties of the related predictor.
 - Extend the approach to the spatio-temporal setting.

References

- E. Bellier *et al.* (2013) Reducing the uncertainty of wildlife population abundance: model-based versus design-based estimates. *Environmetrics*, 24(7):476–488.
- P. Diggle and P. Ribeiro (2007) *Model-based geostatistics*. Springer.
- P. Diggle *et al.* (2013) Spatial and spatio-temporal log-gaussian cox processes: extending the geostatistical paradigm. *Statistical Science*, 28(4):542–563.
- E. Gabriel *et al.* (2016) Adapted kriging to predict the intensity of partially observed point process data. To appear in *Spatial statistics*.
- P. Monestiez *et al.* (2006) Geostatistical modelling of spatial distribution of balaenoptera physalus in the northwestern mediterranean sea from sparse count data and heterogeneous observation efforts, *Ecological Modelling*, 193:615–628.
- A. Tscheschel and D. Stoyan (2006) Statistical reconstruction of random point patterns. *Computational Statistics and Data Analysis*, 51:859–871.
- M-C. van Lieshout and A. Baddeley (2001) Extrapolating and interpolating spatial patterns, *In Spatial Cluster Modelling*, A.B. Lawson and D.G.T. Denison (EDS.) Boca Raton: Chapman And Hall/CRC, pp 61–86.